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Estimating distributions from samples of random size.

Let $\mathbf{X}^{(n)} = (X_1, X_2, \dots, X_n)$ be independent and identically distributed random variables with an unknown distribution function (df) $F(x)$. The distribution of the Kolmogorov statistics (see [1], p.134-141) is based on the empirical df $F_n(x) = \frac{1}{n} \sum_{j=1}^n I(X_j < x)$, where $I(X_j < x)$ is an indicator of the event $(X_j < x)$, over the sample has a *fixed and a priori known size* n . The report considers the problem of estimating the df $F(x)$ and the asymptotic behavior of $F_n(x)$ using a complete (or censored) observations of **a random sizes** ν_p . The random variable ν_p has λ -negative binomial (**NB**) distribution that is the probability mass function of the λ -**NBD** is $\mathbf{P}(\nu_p = 0) = (1 - \lambda q)^\mu$,

$$\mathbf{P}(\nu_p = k) = \frac{a(a + \lambda) \cdot \dots \cdot (a + \lambda(k - a))}{\Gamma(k - a + 2)} (1 - \lambda q)^\mu q^{k-a}, \quad k \in \{a, a + 1, \dots\},$$

$$q = 1 - p, \quad 0 < p, \lambda < 1, \quad \mu = a/\lambda, \quad \nu_p \text{ and } X^{(n)} \text{ are independent.}$$

Let the segment $[0,1]$ be the support of this distribution, $u_i = i/n$, $i = 0, 1, \dots, n$, be the division points on $[0,1]$ and let $W_i = I(X_i < u_i)$. It is also required to construct estimates for the distribution function $F(x)$ and its quantiles $F^{-1}(\alpha)$, $0 < \alpha < 1$, based on the sample $\mathbf{W}^{(n)} = \{(u_i, W_i), i = 0, 1, \dots, n\}$. If the sample size n is *not random* then the kernel-type Nadaraya-Watson estimator $\tilde{F}_n(x)$ with the kernel $A = A(x)$, $\text{supp}(A) = [-1, 1]$, and $\|A\|^2 = \int_{-1}^1 A^2(x) dx < \infty$, was used as estimator of the distribution function $F(x)$ in [2], [3]. There the asymptotic distribution, at $n \rightarrow \infty$, of the normalized difference $n^{2/5}(\tilde{F}_n(x) - \mathbf{E}(\tilde{F}_n(x)))/b(x)$, where $b^2(x) = F(x)(1 - F(x)) \|A\|^2$, was normal $\xi \in \mathcal{N}(0, 1)$. Let $K_\mu(x)$ be the Macdonald function (modified Bessel function of the second kind). Here are results of just a few of them.

Theorem 1. *If $F(x)$ is a continuous distribution function, then*

$$(i) F_{\nu_p}(x) \xrightarrow[\lambda q \rightarrow 1]{p} F(x), \quad (ii) \eta_\lambda = \sqrt{1 - \lambda q} \cdot \frac{\nu_p(F_{\nu_p}(x) - F(x))}{b(x)} \xrightarrow[\lambda q \rightarrow 1]{d} \zeta_\mu, \quad (iii) \eta_\lambda \xrightarrow[\lambda \rightarrow 0]{d} \xi,$$

ζ_μ has a characteristic function $\varphi_\mu(t) = \left(\frac{2}{2 + t^2}\right)^\mu$, and the random variable $\frac{\zeta_\mu}{\sqrt{2}}$ has a distribution density

$$f_\mu(x) = \frac{1}{\Gamma(\mu)\sqrt{\pi}} \left(\frac{|x|}{2}\right)^{\mu-1/2} K_{\mu-1/2}(|x|), \quad x \in \mathbf{R}. \quad (1)$$

The behavior of the kernel estimators and its quantiles based on the samples with random sizes is also considered.

REFERENCES

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 [3] Okumura H., Naito K. *Non-parametric kernel regression for multinomial data. Journal Multivariate Analysis,* **97** (2006), pp. 2009-2022.